Entropy Conserving Probability Transforms and the Entailment Principle

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ABSTRACT

Our main result here is the development of a general procedure for transforming some initial probability distribution into a new probability distribution in a way that the resulting distribution has entropy at least as great as the original distribution. A significant aspect of our approach is that it makes use of the Zadeh's entailment principle which is itself a general procedure for going from an initial possibility distribution to a new possibility distribution so that the resulting possibility has an uncertainty at least as great of the original.

1. Introduction

In [1] Zadeh introduced a general framework for reasoning with uncertain information which he denoted as GTU, for generalized theory of uncertainty. This approach is based on an extension and generalization of his theory of approximate reasoning (AR) [2]. Because of this at times we shall find it convenient to synonymously refer to this as a generalized theory of approximate reasoning (GTAR). Fundamental to this approach is the idea that information can be viewed as a constraint on a variable or collection of variables. In this framework our knowledge base consists of a collection of constraints. A basic reasoning mechanism in the GTU involves the conjunction of constraints in the knowledge base which induces further constraints. Finally inferences are made using these induced constraints. Two other components of the reasoning mechanism in GTU are the use of Zadeh's extension and entailment principles [3].

As indicated in [1, 4] a generalized constraint is the form V *isr* R. Here V is a variable (or joint variable) taking its value in the space X. R is a constraining relation and r is an indexing variable which identifies the modality of the constraint. In [1, 4] Zadeh lists a number of constraint modalities. Among the principle constraint modalities are possibilistic (r = blank), probabilistic (r = blank) and veristic (r = v).

In the case of probabilistic constraints (r = p) R is essentially a probability distribution **P**

over the space X such that p_i is the probability that $V = x_i$. Here of course we require $\sum_i p_i = 1$.

In the case of a possibilistic constraint R is a possibility distribution over X. In this situation $R(x_i) \in [0, 1]$ is the possibility that $V = x_i$. It is often assumed that there exists some x^* such $R(x^*) = 1$. This is called normality. Typically a possibility distribution is generated from a fuzzy set F which is used to precisiate a linguistically expressed value of the variable V [5]. In this situation the possibility of x_i , $R(x_i) = F(x_i)$ the membership grade of x_i in F. In the light of this we shall use the terms fuzzy set and possibility distribution interchangeably.

In the case of veristic constraint V *isv* R, the set R is a fuzzy set corresponding to the set values taken by V. Here V is a value that can take multiple vales. We refer the reader to [6] for a detailed discussion of veristic variables.

The development of the theory of approximate reasoning by Zadeh and others focused mainly on the possibilistic type of constraints and as such a considerable body of literature exists on the manipulation and management of these types of constraints.[7, 8]. While probability theory is highly developed the techniques for managing probabilistic types of constraints within the spirit of approximate reasoning are not as fully developed. Our goal here is to begin to develop some tools for managing probabilistically constrained variables within this generalized theory of approximate reasoning. In particular we will introduce a general procedure for transforming a probability distribution into another probabilistic distribution so that the resulting probability distribution always has at least as much entropy as the original probability distribution. As we shall see this is closely related to Zadeh's entailment principle [2] for possibility distributions.

2. Measuring Information and Uncertainty

In probability theory a well known concept is the entropy. If $\bf P$ is a probability distribution on $X = \{x_1, ..., x_n\}$ where p_i is the probability of x_i then the entropy of $\bf P$ is expressed as

$$H(\mathbf{P}) = -\sum_{i} p_{i} \ln(P_{i})$$

While there exists other formalizations of the concept of entropy [9], this expression, called the Shannon entropy measure, is the most wildly used. It is well known that entropy measures the

uncertainty associated with the probability distribution **P**. Conversely it measures the information contained in **P**. We note that increasing entropy corresponds to more uncertainty, less information about the value of the variable.

An important paradigm associated with the concept of probabilistic entropy is the principle of maximal entropy [10]. This principle has many applications in modern technology. One application is to use it to select from among a number of possible probability distributions. In using this principle we are selecting the distribution with the least information, the most uncertainty.

A related concept within the framework of possibilistic uncertainty is the idea of specificity which has been studied in considerable detail by Yager [11-13] and Klir [14]. While a number of measures for specificity have been suggested we shall find the following one, introduced in [15], to be the most useful for our purposes. Assume F is a possibility distribution on the space $X = \{x_1, ..., x_n\}$. Without loss of generality we shall assume the elements in X have been indexed such that x_1 has the largest membership grade in X. Our measure of specificity of F is

$$Sp(F) = F(x_1) - \frac{1}{n-1} \sum_{j=2}^{n} F(x_j).$$

In effect, Sp(F) is the largest possibility grade in F minus the average of possibility grades of the other elements. We can easily show that $Sp(F) \in [0, 1]$. We also note that Sp(F) uniquely attains its maximal value of 1 for the case when $F(x_1) = 1$ and $F(x_j) = 0$ for all $j \ne 1$. We see Sp(F) attains its minimal value of zero when all elements have the same possibility.

In essence the measure of specificity is measuring the certainty with which we know the value of V based on the constraint V *is* F. It provides a measure of the information contained in the constraining fuzzy set.

We note that decreasing specificity corresponds to more uncertainty, less information about the value of variable. In [16] Dubois and Prade investigate the principle of minimal specificity, a concept analogous to the principle of maximal entropy in probability theory.

A very important special case of F is the normal case, where at least one element has membership grade 1 with our special indexing we have $F(x_1) = 1$; in this case

$$Sp(F) = 1 - \frac{1}{n-1} \sum_{j=2}^{n} F(x_j).$$

We note this special case is closely related to Klir's concept of non-specificity [17].

We note that if F and E are two **normal** possibility distributions (fuzzy sets) such that $F \subseteq E$, $F(x_j) \le E(x_j)$ for all j, then $Sp(F) \ge Sp(E)$. Thus if F and E are two normal fuzzy subsets with F contained in E then the specificity of F is at least as great as E.

3. The Entailment Principle

Within the framework of the theory approximate reasoning, when confined to possibilistic variables, a fundamental role is played by the entailment principle. This principle essentially formalizes the fact that if we know that the value of a variable V lies in the set A then we can naturally say it lies in the set B. Formally we express this principle as follows. If A and B are two fuzzy subsets of X such that $A \subseteq B$, $A(x) \le B(x)$ for all x, then from the proposition V *is* A we can infer the proposition V *is* B. The use of this entailment principle involves a change of an initial possibility distribution to a new possibility such that the information in the resulting possibility distribution is not greater then in the original possibility distribution. In particular if from V *is* A we infer V *is* B where $A \subseteq B$ then $Sp(A) \ge Sp(B)$. Thus we see that in applying the entailment principle we are going from a situation of more specificity (information) to less specificity (information). Essentially we are going from more certainty to less certainty about V.

While the introduction, within the framework of GTU, of a principle for probabilistic constraints analogous to the entailment principle for the case of possibilistic constraints would be a very useful tool our goal here is slightly less ambitious. Our purpose here is only to try to begin the process of providing an analogous entailment principle by looking at the issue of entropy related probability distributions. It would appear that any form of entailment principle for probabilistic constraints should have as one of its features not increasing the information in the original probability distribution. In particular if P is a probability distribution and we induce, using some form of entailment principle, another probability distribution Q then we should require that the

entropies of these two probability distributions satisfy the relation, $H(P) \le H(Q)$. That is Q has as least as large entropy, no more information then P. With this in mind we shall propose a general methodology for transforming a given probability distribution into another probability distribution in such a way that our resulting probability distribution has at least as great entropy as the original. An interesting aspect of our approach is that it makes use of the possibilistic entailment principle.

The basic steps in our methodology are the following. We shall first appropriately translate the probabilistic constraint V isp P into a possibilistic constraint, V is F. We then apply the possibilistic entailment principle on this fuzzy constraint, to give us V is E. We then appropriately retranslate the fuzzy constraint V is E into a probabilistic constraint, V isp Q. As we shall show our methodology will satisfy the property of going for more information to less information. In particular we shall see that the entropies satisfy $H(Q) \ge H(P)$.

4. Possibilistic Probability Transformation

Central to our approach is the use of a probability-possibility transformation. A number of approaches have been suggested in the literature [18-20]. We shall use an approach to possibility-probability transformation which was initially described by Dubois and Prade in [21, 22].

Assume ${\bf P}$ is a probability distribution on $X=\{x_1,...,x_n\}$ where $p_1\geq p_2\geq\geq p_n$. The elements have been indexed in descending order of their probabilities. We associate with this a possibility distribution on X such that u_i is the possibility of x_i where

$$u_n = n p_n$$

$$u_j = j (p_j - p_{j-1}) + u_{j+1}$$
 for $j = n - 1$ to 1 (I)

Example: Consider a probability distribution on $X = \{x_1, x_2, x_3, x_4, x_5\}$ where

$$p_1=0.4,\,p_2=0.3,\,p_3=0.2,\,p_4=0.1,\,p_5=0.$$

In this case we get: $u_5 = 0$, $u_4 = 0.4$, $u_3 = 0.7$, $u_2 = 0.9$, $u_1 = 1$.

Observation: If we let $p_{n+1} = 0$ we can more succinctly express (I) as $u_j = \sum_{k=j}^{n} k(p_k - p_{k+1})$.

From this we see
$$u_j = j p_j + \sum_{k=j+1}^{n} p_k$$

We note some properties of this approach.

Property 1: 1. If
$$p_j = p_{j+1}$$
 then $u_j = u_{j+1}$
2. If $p_j > p_{j+1}$ then $u_j > u_{j+1}$

Property 2: If $p_i = 0$ then $u_i = 0$

Property 3 If $p_j = \frac{1}{n}$ for all j then $u_j = 1$ for all j.

Property 4: It is always the case that $u_1 = 1$

Using formula (I) we can go in the opposite way and get the transformation from a possibility distribution to a probability distribution. Assume $u_1 \ge u_2 \ge ... \ge u_n$ is a normal possibility distribution on X, $u_1 = 1$. We obtain an associated probability distribution on X where

$$p_n = \frac{u_n}{n}$$
 $p_j = p_{j+1} + \frac{u_j - u_{j+1}}{i}$ (II)

We easily see $p_j \ge 0$ for j. Furthermore we also can show the following properties

1) If
$$u_j = u_k$$
 then $p_j = p_k$
If $u_j > u_k$ then $p_j > p_k$

2) If
$$u_{1} = 0$$
 then $p_{1} = 0$

3) If
$$u_j = 1$$
 for all j then $p_j = \frac{1}{n}$ for all j

4)
$$\sum_{j} p_{j} = 1$$

If we denote $u_{n+1}=0$ then we can more succinctly express II as $p_j=\sum_{k=j}^n\frac{u_k-u_{k+1}}{k}$. It also can shown that II can be expressed as $p_j=\frac{u_j}{j}+\sum_{k=j+1}^n(\frac{u_k}{k(k-1)})$.

If the application of I on a probability distribution P leads to a possibility distribution A then the application of II on A brings us back to a probability distribution which is exactly P.

5. Entropy Conserving Probability Transformation

In the following we propose a general method for transforming an initial probability

distribution into another probability distribution in such a way that the resulting probability distribution always has at least as much entropy as the original probability distribution.

Assume P is a probability distribution on $X = \{x_1, ..., x_n\}$ indexed such that $p_1 \ge p_2 \ge \ge p_n$. We define a **ET** transformation of the probability distribution P into a new probability distribution Q, **ET**(P) —> Q as follows:

- 1. We use the probability-possibility transform I on P to induce a possibility distribution A on X such that the possibility of x_i is a_i where $a_i = i$ $p_i + \sum_{k=i+1}^{n} p_k$
- **2.** Apply the possibilistic entailment principle on A to generate the possibility distribution B. Specifically, we apply one of the available transformations of A implementing the entailment principle. This results in a new possibility distribution B such that $b_i \ge a_i$ for all i.
- **3.** Finally, after reordering if necessary, we apply the possibility to probability transformation II on B to obtain a probability distribution Q on X, $q_i = \sum_{k=1}^n \frac{b_k b_{k+1}}{k}$.

The following theorem provides a very significant relation between the entropy of the original probability distribution and the resulting probability under this ET transformation

Theorem: If ET(P) = Q then $H(P) \le H(Q)$

Proof: We start with P and using I we induce the possibility distribution A. We can denote these as $(p_1, ..., p_n)$ and $(a_1, ..., a_n)$. We obtain B from b_j by adding some value to a_j hence $b_j \ge a_j$. For simplicity we shall let $b_{(i)}$ be a permutation of the b_i such $b_{(i)}$ is the i^{th} largest of the elements in B. We can denote $B = (b_{(1)}, b_{(2)}, ..., b_{(n)}) = (g_1, g_2, ..., g_n)$. It is easy to see that for each i $b_{(i)} \ge a_i$. We shall let F denote the process of going from a possibility distribution to a probability distribution. Thus F(A) takes A into a probability distribution and F(B) takes B into a probability distribution. Furthermore F(A) = P our original probability distribution. Our objective is to show that $H(F(B)) \ge H(F(A))$ that is the entropy of the probability distribution induced by B is at least as great as the entropy of P under the condition $b_{(i)} \ge a_i = a_{(i)}$.

In order to prove this it is sufficient to show that if we start with $A=(a_1,...,a_n)$ and increase any element in A to obtain A' then $H(F(A')) \geq H(F(A))$. More directly this simply requires us to prove that $\frac{\partial H(F(A))}{\partial a_{l'}} \geq 0$

By applying the chain rule formula to the expression $H(F(A)) = \sum_{j=1}^{n} -p_{j} \log p_{j}$ we get

$$\frac{\partial H(F(A))}{\partial a_k} = \sum_{i=1}^{n} (-\log(p_i) - 1) \frac{\partial p_i}{\partial a_k}$$

From this we get

$$\frac{\partial H(F(A))}{\partial a_k} = -\sum_{i=1}^n \log(p_i) \frac{\partial p_i}{\partial a_k} - \sum_{i=1}^n \frac{\partial p_i}{\partial a_k}$$

Since $\sum_{i=1}^{n} p_i = 1$, we can conclude that

$$0 = \frac{\partial (\sum_{i=1}^{n} p_i)}{\partial a_k} = \sum_{i=1}^{n} \frac{\partial p_i}{\partial a_k}$$

This implies

$$\frac{\partial H(F(A))}{\partial a_k} = -\sum_{i=1}^{n} \log (P_i) \frac{\partial P_i}{\partial a_k}$$

Due to the fact that

$$p_i = \frac{1}{i} a_i - \sum_{j=i+1}^{n} \frac{1}{(j-1)j} a_j$$

we obtain

$$\begin{split} \frac{\partial p_i}{\partial a_k} &= - \ \frac{1}{(k-1)(k)} & \qquad \text{for } i < k \\ \frac{\partial p_k}{\partial a_k} &= \frac{1}{k} & \qquad \text{for } i = k \\ \frac{\partial p_i}{\partial a_k} &= 0 & \qquad \text{for } i > k \end{split}$$

Substituting these into $\frac{\partial H(F(A))}{\partial a_k} = -\sum_{i=1}^n \log(p_i) \frac{\partial p_i}{\partial a_k}$ we get

$$\frac{\partial H(F(A))}{\partial a_k} = -\frac{1}{k}log(p_k) + \frac{1}{(k)(k-1)}\sum_{i=1}^{k-1}log(p_i)$$

since $p_1 \ge p_2 \ge p_3 \ge ... \ge p_n$ then for i < k we have $p_i \ge p_k$ and hence $log(p_i) \ge log(p_k)$.

From this we get

$$\begin{split} &\frac{\partial H(F(A))}{\partial a_k} \geq -\frac{1}{k}log(p_k) + \frac{1}{(k)(k-1)}\sum_{i=1}^{k-1}log(p_k) \\ &\frac{\partial H(F(A))}{\partial a_k} \geq -\frac{1}{k}log(p_k) + (k-1)\frac{1}{(k-1)}\frac{1}{k}log(p_k) \\ &\frac{\partial H(F(A))}{\partial a_k} \geq -\frac{1}{k}log(p_k) + \frac{1}{k}log(p_k) \geq 0 \end{split}$$

Thus we see that applying the ET transformation on a probability distribution P always results in a probability distribution with more entropy, it tends to increase the uncertainty.

We observe that for any probability distribution P there always exists an ET transformation into the probability distribution Q such that $q_i = \frac{1}{n}$ for all i. We see this as follows. If from P we get the possibility distribution A with values a_i then if we increase each a_i by Δ_i such that $b_i = a_i + \Delta_i = 1$ then in this case $q_i = \frac{1}{n}$ for all i..

Another important property of the ET transformation is the following.

Property: Let P be a probability distribution such that p_1 , the probability of x_1 , is the largest. Then if ET(P) = Q we always have that q_1 , the probability of x_1 , is also always the largest, $q_1 \ge q_i$ for all j.

We shall say that an ET transform is order preserving if ordering of $p_j \ge p_k$ results in $q_j \ge q_k$. We can guarantee this condition as follows. Assume $p_j \ge p_k$ for j < k and let a_j be the possibility transformation of p_j . In this case $a_1 \ge a_2, ..., \ge a_n$. If we modify a_j to b_j , ie $b_j = a_j + \Delta_j$ such that $b_1 \ge b_2, ..., \ge b_n$ then $q_1 \ge q_2, ..., \ge q_n$. Here we get order preservation.

It is interesting to observe the effect of modifying one of the a_i . Again assume we start with P where $p_1 \ge p_2 \ge ... \ge p_n$. From this we generate the possibility distribution $a_1 \ge a_2 \ge ... \ge a_n$. If we directly use the a_i to obtain q_i we get $q_j = \sum_{k=j}^n \frac{a_k - a_{k+1}}{k} = p_j$. Assume we just modify a_i by adding Δ , thus $b_i = a_i + \Delta$ and $b_j = a_j$ for all $j \ne i$. Using the fact that $q_j = \sum_{k=j}^n \frac{b_k - b_{k+1}}{k}$. First we see that for j > i, the smaller elements, since $a_k = b_k$ for k > i we get

$$q_j = \sum_{k=j}^n \frac{b_k - b_{k+1}}{k} = \sum_{k=j}^n \frac{a_k - a_{k+1}}{k} = p_j$$

For j = i we have

$$q_{i} = \sum_{k=j}^{n} \frac{b_{k} - b_{k+1}}{k} = \frac{b_{i} - b_{i+1}}{i} + \sum_{k=i+1}^{n} \frac{b_{k} - b_{k+1}}{k}$$

Since for k > i, $b_k = a_i$ and $b_i = a_i + \Delta$ we

$$q_{i} = \frac{a_{i} + \Delta - a_{i+1}}{i} + \sum_{k=i+1}^{n} \frac{a_{k} - a_{k+1}}{k}$$

$$q_{i} = \frac{\Delta}{i} + \sum_{k=i}^{n} \frac{a_{k} - a_{k+1}}{k}$$

$$q_{i} = \frac{\Delta}{i} + p_{i}$$

Thus the ith largest probability has increased by $\frac{\Delta}{i}$.

Consider now any j < i, the elements with larger probabilities. In this case again,

$$\begin{split} q_j &= \sum_{k=j}^n \frac{b_k - b_{k+1}}{k} \\ q_j &= \sum_{k=j}^{i-2} \frac{b_k - b_{k+1}}{k} + \frac{(b_{i-1} - b_i)}{i-1} + \frac{(b_i - b_{i-1})}{i} + \sum_{k=i+1}^n \frac{b_k - b_{k+1}}{k} \\ q_j &= \sum_{k=j}^{i-2} \frac{a_k - a_{k+1}}{k} + \frac{(a_{i-1} - (a_i + \Delta))}{i-1} + \frac{((a_i + \Delta) - a_{i-1})}{i} + \sum_{k=i+1}^n \frac{a_k - a_{k+1}}{k} \\ q_j &= \sum_{k=j}^n \frac{a_k - a_{k+1}}{k} - \frac{\Delta}{i-1} + \frac{\Delta}{i} = P_j - \frac{\Delta}{(i-1)i} \\ q_j &= P_j - (\frac{\Delta}{i}) \frac{1}{(i-1)} \end{split}$$

We see that the amount added to p_i , $\frac{\Delta}{i}$, has been accounted for by uniformly subtracting $(\frac{1}{i-1})\frac{\Delta}{i}$ from all probabilities greater than p_i .

Thus we see in this case if we just increase a_i by Δ the following changes in P have happened. All probabilities less the p_i have been unchanged. p_i has been increased by $\frac{\Delta}{i}$. All probabilities greater then p_i have been diminished by $(\frac{\Delta}{i})\frac{1}{(i-1)}$.

Let us now consider the case in which we modify two of the possibility distribution. For simplicity we consider changing two contiguous ones. Thus here we let $b_i = a_i + \Delta$ and $b_{i-1} = a_{i-1} + d$ and $b_j = a_j$ for all $j \neq i$, i-1. Here again $q_j = \sum_{k=j}^n \frac{b_k - b_{k+1}}{k}$. From this we easily see that for j > i, $q_j = p_j$ no change. For j = i, $q_i = p_i + \frac{\Delta}{i}$. This is as in the preceding. Consider now j = i-1. Here we have

$$\begin{aligned} q_{i-1} &= \sum_{k = i-1}^{n} \frac{b_k - b_{k+1}}{k} = \frac{b_{i-1} - b_i}{i-1} + \frac{b_i - b_{i+1}}{i} + \sum_{k = i+1}^{n} \frac{b_k - b_{k+1}}{k} \\ q_{i-1} &= = \frac{a_{i-1} + d - (a_i + \Delta)}{i-1} + \frac{a_i + \Delta - a_{i+1}}{i} + \sum_{k = i+1}^{n} \frac{b_k - b_{k+1}}{k} \end{aligned}$$

$$q_{i-1} = \frac{d - \Delta}{i - 1} + \frac{\Delta}{i} + \sum_{R=i-1}^{n} \frac{a_k - a_{k+1}}{k}$$

$$q_{i-1} = p_{i-1} + \frac{d}{i-1} - \frac{\Delta}{(i)(i-1)}$$

We can easily show that for j < i - 1 we have

$$q_j = p_j - \frac{\Delta}{(i)(i-1)} - \frac{d}{(i-1)(i-2)}$$
.

Based upon the preceding we can make some general observations about the relationship of q_j and p_j for the order preserving case. Let $b_j = a_j + \Delta_j$ for j = 2 to n. Note we can't change a_1 as it already equals 1. In this case we have that $q_j = p_j + \frac{\Delta_j}{j} - \sum_{k=j-1}^n \frac{\Delta_k}{(k)(k-1)}$. We observe that the smallest probability to increase its value of a_j , have a non-zero Δ_j , will always display an increase in probability. We also note that since Δ_1 must always be zero then if any of the a_i increases we will have a decrease in p_1 , that is $q_1 < p_1$.

We not something similar was done in Dubois and Hullermeier [23] where the authors used the possibility probability transformation proposed by Moral [24].

6. Certainty Qualification of Probabilistic Constraints

We now consider a special case of ET transformations. Here we let $\Delta_j=(1$ - $a_j)$ α where $\alpha \in [0,1]$. Thus we add an amount proportional to the distance from 1. Here then we have for k=1 to n that

$$b_k = a_k + (1 - a_k)\alpha$$

$$b_K = \alpha + (1 - \alpha) \ a_K$$

We still require that $b_{n+1} = a_{n+1} = 0$. In order to guarantee this we can express

$$b_{n+1} = \alpha + (1 - \alpha)a_{n+1} - \alpha$$

Using this we have

$$\begin{split} q_j &= \sum_{k \,=\, 1}^n \,\, \frac{b_k \, -\, b_{k \,+\, 1}}{k} = \sum_{k \,=\, 1}^{n \, -\, 1} \, \frac{b_k \, -\, b_{k \,+\, 1}}{k} \, + \frac{b_n \, -\, b_{n \,+\, 1}}{n} \\ q_j &= \sum_{k \,=\, 1}^{n \, -\, 1} \,\, \frac{\alpha \, +\, (1 \, -\, \alpha) a_k \, -\, (\alpha \, +\, (1 \, -\, \alpha) a_{k+1})}{k} \, + \frac{\alpha \, +\, (1 \, -\, \alpha) a_n}{n} \, - \frac{(\alpha \, +\, (1 \, -\, \alpha) a_{n \,+\, 1} \, -\, \alpha)}{n} \end{split}$$

$$\begin{split} q_j &= \sum_{k=1}^n \frac{\alpha + (1 - \alpha) a_k - (\alpha + (1 - \alpha) a_{k+1})}{k} + \frac{\alpha}{n} \\ q_j &= (1 - \alpha) \sum_{k=1}^n \frac{a_k - a_{k+1}}{k} + \frac{\alpha}{n} \\ q_j &= (1 - \alpha) \ p_j + \alpha \ \frac{1}{n} \end{split}$$

Thus here as α goes from zero to one we move from our given probability distribution to a state of complete ignorance.

This transformation can have an interesting use within the framework of the theory of Generalized Approximate reasoning. Assume we have a possibilistic constraint V *is* G with which we have an associated degree of confidence λ . We recall one approach [25, 26] to representing this is to discount the basic statement by using a constraint V *is* H where $H(x_i) = Max[G(x_i), 1 - \lambda]$. This process is called certainty qualification.

The preceding ET transformation can provide the basis for an analogous operator of certainty qualification for a probabilistic constraint. Assumes we have a probabilistic constraint V *isp* P with which we associate a of degree of certainty or confidence λ . We can transform this to another probabilistic constraint V *isp* Q where

$$Q(x_i) = \lambda p(x_i) + (1 - \lambda) \frac{1}{n}$$

In this we see $\alpha = 1 - \lambda$. Thus as our confidence in the original information decreases, λ , goes to zero we get close to completely discounting the distribution provided.

7. Conclusion

Our main result here is the development of a general procedure for transforming some initial probability distribution into a new probability distribution in a way that the resulting distribution has entropy at least as great as the original distribution. An significant aspect of our approach is that it makes use of the Zadeh's entailment principle which is itself a general procedure for going from an initial possibility distribution to a new possibility distribution so that the resulting possibility has an uncertainty at least as great of the original.

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