

# Confirmation Bias in Systems Engineering: A Pedagogical Example

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## Abstract

One of the biases potentially affecting systems engineers is the confirmation bias, when instead of selecting the best hypothesis based on the data, people stick to the previously-selected hypothesis until it is disproved. In this paper, on a simple example, we show how important is to take care of this bias: namely, that because of this bias, we need twice as many experiments to switch to a better hypothesis.

## 1 Formulation of the Problem

**Confirmation bias.** It is known that our intuitive reasoning shows a lot of unexpected biases; see, e.g., [2]. One of such biases is a *confirmation bias*, when, instead of selecting the best hypothesis based on the data, people stick to the previously-selected hypothesis until it is disproved. This bias is ubiquitous in systems engineering; see, e.g., [1, 4, 5, 8].

**How important is it to take the confirmation bias into account?** Taking care of the confirmation bias requires some extra effort; see, e.g., [3, 7, 8, 9] and references therein. A natural question is: is the resulting improvement worth this extra effort? How better the result will we get?

In this paper, on a simple example, we show that the result is drastically better: namely, that if we properly take this bias into account, then we will need half as many experiments to switch to a more adequate hypothesis.

## 2 Analysis of the Problem

**Description of the simple example.** Let us consider the simplest possible case when we have a parameter  $a$  that may be 0 and may be non-zero, and we directly observe this parameter. We will also make the usual assumption that the observation inaccuracy is normally distributed, with 0 mean and known standard deviation  $\sigma$ .

In this case, what we observe are the values  $x_1, \dots, x_n$  which are related to the actual (unknown) value  $a$  by a relation  $x_i = a + \varepsilon_i$  ( $i = 1, \dots, n$ ), where  $\varepsilon_i$  are independent normally distributed random variables with 0 means and standard deviation  $\sigma$ .

**Two approaches.** In the ideal approach, we select one of the two models – the null-hypothesis  $a = 0$  or the alternative hypothesis  $a \neq 0$  – by using the usual Akaike Information Criterion (AIC); see, e.g., [6].

In the confirmation-bias approach, we estimate the value  $a$  based on the observations  $x_1, \dots, x_n$ , and we select the alternative hypothesis only if the resulting estimate is statistically significantly different from 0 – i.e., e.g., that the 95% confidence interval for the value  $a$  does not contain 0.

**What if we use AIC.** In the AIC, we select a model for which the difference  $\text{AIC} \stackrel{\text{def}}{=} 2k - 2 \ln(\widehat{L})$  is the smallest, where  $k$  is the number of parameters in a model and  $\widehat{L}$  is the largest value of the likelihood function  $L$  corresponding to this model.

The null-model  $a = 0$  has no parameters at all, so for this model, we have  $k = 0$ . For  $n$  independent measurement results, the likelihood function is equal to the product of the values

$$\frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{x_i^2}{2\sigma^2}\right)$$

of the Gaussian probability density function corresponding to these measurement results  $x_i$ . Thus,

$$L = \prod_{i=1}^n \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{x_i^2}{2\sigma^2}\right)$$

and so, for this model,

$$\text{AIC}_0 = -2 \ln(L) = 2n \cdot \ln(\sqrt{2\pi} \cdot \sigma) + \frac{1}{\sigma^2} \cdot \sum_{i=1}^n x_i^2.$$

We assume that  $x_i = a + \varepsilon_i$ , where the mean value of  $\varepsilon_i$  is 0 and the standard deviation is  $\sigma$ . Thus, the expected value of  $x_i^2$  is equal to  $a^2 + \sigma^2$ . For large values  $n$ , due to the Law of Large Numbers (see, e.g., [6]), the average  $\frac{1}{n} \cdot \sum_{i=1}^n x_i^2$  is approximately equal to the expected value  $E[x_i^2] = a^2 + \sigma^2$ . Thus,

$\sum_{i=1}^n x_i^2 \approx n \cdot (a^2 + \sigma^2)$  and hence,

$$\text{AIC}_0 = 2n \cdot \ln\left(\sqrt{2\pi} \cdot \sigma\right) + \frac{1}{\sigma^2} \cdot n \cdot (a^2 + \sigma^2). \quad (1)$$

The alternative model  $a \neq 0$  has one parameter  $a$ , so here  $k = 1$ . The corresponding likelihood function is then equal to

$$L = \prod_{i=1}^n \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{(x_i - \hat{a})^2}{2\sigma^2}\right).$$

We select the parameter  $a$  that maximizes the value of this likelihood function. Maximal likelihood is the usual way of estimating the parameters, which in this case leads to  $\hat{a} = \frac{1}{n} \cdot \sum_{i=1}^n x_i$ . For large  $n$ , this estimate is close to the actual value  $a$ , so we have

$$\hat{L} = \prod_{i=1}^n \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{(x_i - a)^2}{2\sigma^2}\right).$$

For this model,  $x_i - a = \varepsilon_i$ , thus,

$$\text{AIC}_1 = 2 - 2 \ln(\hat{L}) = 2 + 2n \cdot \ln\left(\sqrt{2\pi} \cdot \sigma\right) + \frac{1}{\sigma^2} \cdot \sum_{i=1}^n \varepsilon_i^2.$$

For large  $n$ , we have  $\sum_{i=1}^n \varepsilon_i^2 \approx n \cdot \sigma^2$ , hence

$$\text{AIC}_1 = 2 + 2n \cdot \ln\left(\sqrt{2\pi} \cdot \sigma\right) + \frac{1}{\sigma^2} \cdot n \cdot \sigma^2. \quad (2)$$

The second model is preferable if  $\text{AIC}_1 < \text{AIC}_0$ . By deleting common terms in these two values  $\text{AIC}_i$ , we conclude that the desired inequality reduces to  $2 < \frac{n \cdot a^2}{\sigma^2}$ , i.e., equivalently, to

$$n > \frac{2\sigma^2}{a^2}. \quad (3)$$

**What if we use a confirmation-bias approach.** In the confirmation-bias approach, we estimate  $a$  – and we have already mentioned that the optimal estimate is  $a = \frac{1}{n} \cdot \sum_{i=1}^n x_i$ . It is known (see, e.g., [6]) that the standard deviation of this estimate is equal to  $\sigma_e = \frac{\sigma}{\sqrt{n}}$ . Thus, the corresponding 95% confidence interval has the form  $[a - 2\sigma_e, a + 2\sigma_e]$ . The condition that this interval does not

contain 0 is equivalent to  $|a| > 2\sigma_e$ , i.e., equivalently, to  $a^2 > 4\sigma_e^2$ . Substituting the above expression for  $\sigma_e$  into this inequality, we conclude that  $a^2 > 4 \cdot \frac{\sigma^2}{n}$ , i.e., equivalently, that

$$n > \frac{4\sigma^2}{a^2}. \quad (4)$$

**Conclusion.** By comparing the expressions (3) and (4) corresponding to the two approaches, we can indeed see that the confirmation-bias approach requires twice as many measurements than the approach in which we select the best model based on the data.

Thus indeed, avoiding confirmation bias can lead to a drastic improvement in our estimates and thus, in our decisions.

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