# Adding Constraints to Situations When, In Addition to Intervals, We Also Have Partial Information about Probabilities

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Limitations of the . . . How to Describe . . . Kolmogorov-Smirnov . . . Illustration: . . . Illustration: . . . Computing VComputing  $\overline{V}$ Computational . . . How to Handle Gauging Amount of . . . Case of a Continuous... Case of a p-Box Acknowledgments Shannon's Derivation: . . Shannon's Derivation . . • Page 1 of 19 Go Back Full Screen

#### 1. Statistical Analysis Is Important

- Fact: statistical analysis of measurement and observation results is an important part of data processing and data analysis.
- Specifics:
  - when faced with new data,
  - engineers and scientists usually start with estimating standard statistical characteristics such as:
    - \* the mean E,
    - \* the variance V,
    - \* the probability distribution function (pdf) F(x) of each variable, and
    - \* the covariance and correlation between different variables.



# 2. Limitations of Traditional Statistical Techniques and the Need to Consider Interval Uncertainty

- Main assumption: traditional statistical techniques assume that the measured values  $\tilde{x}_1, \dots \tilde{x}_n$  coincide with the actual values  $x_1, \dots, x_n$  of the measured quantities.
- This assumption is often true: if the variability of each variable is much higher than the measurement errors  $\Delta x_i \stackrel{\text{def}}{=} \widetilde{x}_i x_i$ .
- Example: the accuracy of measuring a person's height ( $\approx 1$  cm) is  $\ll$  variability in height.
- Sometimes, this assumption is not true: when the measurement errors  $\Delta x_i$  are of the same order of magnitude.
- Conclusion:  $\Delta x_i$  cannot be ignored in statistical analysis.
- Frequent situation: the only information about  $\Delta x_i$  is the upper bound  $\Delta_i$ :  $|\Delta x_i| \leq \Delta_i$ .
- Interval uncertainty: the only information about  $x_i$  is that  $x_i \in \mathbf{x}_i \stackrel{\text{def}}{=} [\widetilde{x}_i \Delta_i, \widetilde{x}_i + \Delta_i].$



# 3. Adding Interval Uncertainty to Statistical Techniques: What Is Known

- We start with: a statistic  $C(x_1, \ldots, x_n)$ , such as:
  - population mean  $E = \frac{1}{n} \cdot \sum_{i=1}^{n} x_i;$
  - population variance  $V = \frac{1}{n} \cdot \sum_{i=1}^{n} (x_i E)^2$ ;
  - histogram pdf  $F_n(x) = \frac{\#i : x_i \le x}{n}$ ;
  - population covariance  $C_{x,y} = \frac{1}{n} \cdot \sum_{i=1}^{n} (x_i E_x) \cdot (y_i E_y).$
- Interval extension: find the range

$$\mathbf{C} = C(\mathbf{x}_1, \dots, \mathbf{x}_n) \stackrel{\text{def}}{=} \{C(x_1, \dots, x_n) : x_1 \in \mathbf{x}_1, \dots, x_n \in \mathbf{x}_n\}.$$

- $\bullet$  General case: the general problem is NP-hard, even for V.
- $\bullet$  Conclusion: in general, we can only compute an enclosure.
- Specific cases: efficient algorithms are possible: for  $\mathbf{E}$ , for  $\underline{V}$ , for  $\overline{V}$  when  $[\underline{x}_i, \overline{x}_i] \not\subseteq (\underline{x}_j, \overline{x}_j)$ , etc.

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### 4. Limitations of the Existing Approach

- Currently used idea:
  - we start with a statistic  $C(x_1, \ldots, x_n)$  for estimating a given characteristic S;
  - we evaluate this statistic under interval uncertainty, resulting in  $\mathbf{C} = C(\mathbf{x}_1, \dots, \mathbf{x}_n)$ .
- First limitation of this idea:
  - we know that  $C(x_1, \ldots, x_n) \approx S$ ;
  - sometimes, the estimation error  $C(x_1, \ldots, x_n) S \neq 0$  is not always taken into account when estimating  $\mathbb{C}$ .
- Solution: instead of the original statistic C, we consider the bounds  $C^-$  and  $C^+$  of the confidence interval.
- Good news: the interval  $\left[\underline{C}^-, \overline{C}^+\right]$  is an enclosure for S (with appropriate certainty).
- Remaining limitation: excess width.
- New idea: find  $S = \{S(F) : F \text{ is possible}\}.$
- Related problem: how to describe class  $\mathcal{F}$  of possible probability distributions F.

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# 5. How to Describe Possible Probability Distributions: p-Boxes

- General situation:
  - we do not know the probability distribution of the actual values  $x_i$ ;
  - we want to determine this distribution.
- Question: which characteristics of this distribution are practically useful?
- Practical example:
  - there is a critical threshold  $x_0$  after which a chip delays too much, a panel cracks, etc.;
  - we want to make sure that the probability of exceeding  $x_0$  is small.
- Resulting characteristic:  $Prob(x_i \le x_0)$ , i.e.,  $cdf F(x_0)$ .
- *p-box*:
  - we cannot determine the *exact* values of F(x);
  - thus, we should look for bounds  $\mathbf{F}(x) = [\underline{F}(x), \overline{F}(x)];$
  - the function  $x \to \mathbf{F}(x)$  is called a *p-box*.

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# 6. Kolmogorov-Smirnov (KS) p-Box

- New idea (reminder):
  - transform observations  $x_1, \ldots, x_n$  into a p-box;
  - estimate a characteristic S based on the p-box.
- How to transform: use KS inequalities.
- Main idea behind KS: for each  $x_0$ , we have
  - actual (unknown) probability  $p = F(x_0)$  that  $x \leq x_0$ , and
  - frequency  $F_n(x_0) = \frac{\#i : x_i \le x_0}{n}$ .
- Known: for large n,  $F_n(x_0) \approx \text{normal}$ , and with given certainty  $\alpha$ , we have  $p k \cdot \sigma \leq F_n(x_0) \leq p + k \cdot \sigma$ , where  $\sigma = \sqrt{\frac{p \cdot (1-p)}{n}}$  and  $k = k(\alpha)$ .
- Conclusion: with certainty  $\alpha$ , we get bounds on  $p = F(x_0)$  in terms of  $F_n(x_0)$ .
- We use these bounds for  $x_0 = x_i$  and use monotonicity to get bounds  $[F_n(x) \varepsilon, F_n(x) + \varepsilon]$  for all  $x \in [x_i, x_{i+1}]$ .

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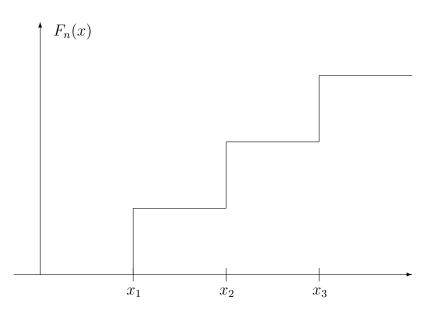
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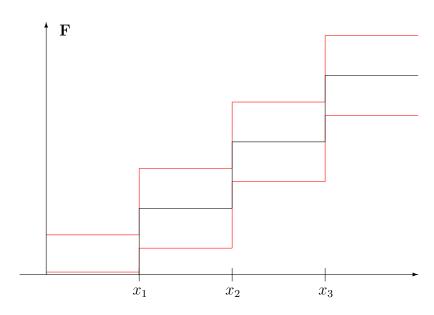
#### 7. Illustration: Histogram Pdf



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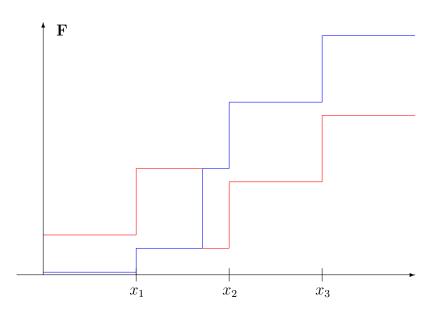
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## 8. Illustration: Kolmogorov-Smirnov p-Box



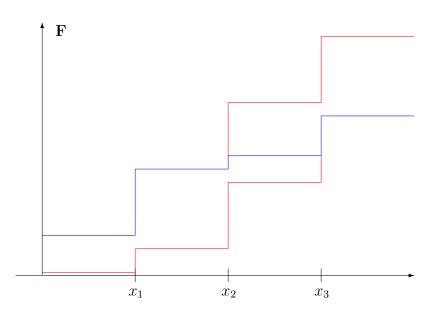
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# 9. Computing $\underline{V}$



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# 10. Computing $\overline{V}$



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# 11. Computational Complexity of Computing $\underline{V}$ and $\overline{V}$

- Traditional method:
  - we can compute V in linear time O(n);
  - computing  $\overline{V}$  is, in general, NP-hard;
  - when  $[\underline{x}_i, \overline{x}_i] \not\subseteq (\underline{x}_i, \overline{x}_i)$ , we can compute  $\overline{V}$  is linear time.
- Analysis:
  - in effect, the variance of  $F \in \mathbf{F}$  can be reduced to the variance over horizontal layers;
  - these layers satisfy the above "subset" property.
- New method:
  - we can compute  $\underline{V}$  in linear time O(n), and
  - we can compute  $\overline{V}$  in linear time O(n);

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#### 12. How to Handle Additional Constraints

- Previously: the only information we have is  $F(x) \in \mathbf{F}(x)$ .
- Frequent situation: we have additional information about F(x).
- Example: we know the shape of F(x), i.e., we know that  $F(x) = F_0(x, a_1, \ldots, a_n)$  for known  $F_0$  and  $a_i \in [\underline{a}_i, \overline{a}_i]$ .
- Typical situation:  $F(x) = F_0\left(\sum_{i=1}^n a_i \cdot e_i(x)\right)$ .
- Example: Gaussian  $F(x) = F_0\left(\frac{x-a}{\sigma}\right) = F_0(a_1 \cdot x + a_2).$
- p-box solution: find a p-box containing all such F(x), and estimate, e.g.,  $\mathbf{V}$ , based on this p-box.
- Drawback: excess width.
- Exact estimates:  $\underline{F}(x_i) \leq F_0\left(\sum_{i=1}^n a_i \cdot e_i(x_i)\right) \leq \overline{F}(x_i)$ , hence

$$F_0^{-1}(\underline{F}(x_i)) \le \sum_{i=1}^n a_i \cdot e_i(x_i) \le F_0^{-1}(\overline{F}(x_i)). \tag{*}$$

• Algorithm: apply linear programming to (\*) and  $\underline{a}_i \leq a_i \leq \overline{a}_i$ .

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## 13. Gauging Amount of Uncertainty

- Shannon's idea: (average) number of "yes"-"no" (binary) questions that we need to ask to determine the object.
- Fact: after q binary questions, we have  $2^q$  possible results.
- Discrete case: if we have n alternatives, we need q questions, where  $2^q \ge n$ , i.e.,  $q \sim \log_2(n)$ .
- Discrete probability distribution:  $q = -\sum p_i \cdot \log_2(p_i)$ .
- Continuous case definition: number of questions to find an object with a given accuracy  $\varepsilon$ .
- Interval uncertainty: if  $x \in [a, b]$ , then  $q \sim S \log_2(\varepsilon)$ , with  $S = \log_2(b a)$ .
- Probabilistic uncertainty:  $S = -\int \rho(x) \cdot \log_2 \rho(x) dx$ .



### 14. Case of a Continuous Probability Distribution

- Once an approximate value r is determined, possible actual values of x form an interval  $[r \varepsilon, r + \varepsilon]$  of width  $2\varepsilon$ .
- So, we divide the real line into intervals  $[x_i, x_{i+1}]$  of width  $2\varepsilon$  and find the interval that contains x.
- The average number of questions is  $S = -\sum p_i \cdot \log_2(p_i)$ , where the probability  $p_i$  that  $x \in [x_i, x_{i+1}]$  is  $p_i \approx 2\varepsilon \cdot \rho(x_i)$ .
- So, for small  $\varepsilon$ , we have

$$S = -\sum \rho(x_i) \cdot \log_2(\rho(x_i)) \cdot 2\varepsilon - \sum \rho(x_i) \cdot 2\varepsilon \cdot \log_2(2\varepsilon),$$

where the first sum in this expression is the integral sum for the integral  $S(\rho) \stackrel{\text{def}}{=} - \int \rho(x) \cdot \log_2(\rho(x)) dx$ , so

$$S \approx -\int \rho(x) \cdot \log_2(\rho(x)) dx - \log_2(2\varepsilon).$$



### 15. Case of a p-Box

• Situation: we know that

$$F(x) \in \mathbf{F}(x) = [F_0(x) - \Delta F(x), F_0(x) + \Delta F(x)],$$

where  $F_0(x)$  is smooth, with  $\rho_0(x) \stackrel{\text{def}}{=} F'_0(x)$ .

- Problem: find the range  $[\underline{S}, \overline{S}] = \{S_{\varepsilon}(F) : F \in \mathbf{F}\}.$
- *Known result:* asymptotically,

$$\overline{S} \sim -\int \rho_0(x) \cdot \log_2(\rho_0(x)) dx - \log_2(2\varepsilon).$$

- New result:  $\underline{S} \sim -\int \rho_0(x) \cdot \log_2(\max(2\Delta F(x), 2\varepsilon \cdot \rho_0(x))) dx$ .
- Comment: when  $\varepsilon \to 0$ ,  $\overline{S} \to \infty$  but  $\underline{S}$  remains finite.
- Idea of the proof:  $p_i \approx \rho_0(x_i) \cdot \Delta x_i$ , hence

$$-\sum p_i \cdot \log_2(p_i) \approx -\int \rho_0(x) \cdot \log(\rho_0(x) \cdot \Delta x) \, dx.$$

$$\overline{F}(x) / \sum_{i=0}^{\infty} P_i(x) = 0$$

Here, 
$$\Delta x_i = \max\left(\frac{2\Delta F(x)}{\rho_0(x)}, 2\varepsilon\right)$$
:  $E(x)$ 

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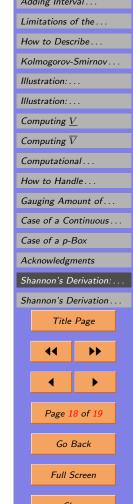


#### 17. Shannon's Derivation: Reminder

- Situation: we know the probabilities  $p_1, \ldots, p_n$  of different alternatives.
- $\bullet$  We repeat the selection N times.
- Let  $N_i$  be number of times when we get  $A_i$ .
- For big N, the value  $N_i$  is  $\approx$  normally distributed with average  $a = p_i \cdot N$  and  $\sigma = \sqrt{p_i \cdot (1 p_i) \cdot N}$ .
- With certainty depending on  $k_0$ , we conclude that

$$N_i \in [a - k_0 \cdot \sigma, a + k_0 \cdot \sigma].$$

- Let  $N_{\text{con}}(N)$  be the number of situations for which  $N_i$  is within these intervals.
- Then, for N repetitions, we need  $q(N) = \log_2(N_{\text{cons}})$  questions.
- Per repetition, we need S = q(N)/N questions.



#### 18. Shannon's Derivation (cont-d)

- Shannon's theorem:  $S \to -\sum p_i \cdot \log_2(p_i)$ .
- Proof:

$$\frac{N!}{N_1!(N-N_1)!} \cdot \frac{N_{\text{cons}} \sim}{N_2!(N-N_1-N_2)!} \cdot \dots = \frac{N!}{N_1!N_2!\dots N_n!}$$

where  $k! \sim (k/e)^k$ . So,

$$N_{\rm cons} \sim \frac{\left(\frac{N}{e}\right)^{N_1}}{\left(\frac{N_1}{e}\right)^{N_1} \cdot \dots \cdot \left(\frac{N_n}{e}\right)^{N_n}}$$

Since  $\sum N_i = N$ , terms  $e^N$  and  $e^{N_i}$  cancel each other.

• Substituting  $N_i = N \cdot f_i$  and taking logarithms, we get  $\log_2(N_{\text{cons}}) \approx -N \cdot f_1 \cdot \log_2(f_1) - \ldots - N \cdot f_n \log_2(f_n).$ 

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